

# Optimal match length in knock-out tournaments

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# Outline

1 Motivation

2 Single match setting

3 Knockout tournament setting

**The goal of a tournament is to single out the best player.**

Some sports play **deterministic rules**. e.g.,

- **Soccer:** 90 minutes regulation + 30 minutes extra time if needed.  
Historically some sequential rule (golden goal).
- **Basketball:** Each game is always 48 minutes.
- **American football:** Each game is always 60 minutes.

And some sports play **sequential rules**. e.g.,

- **Table tennis:** Game ends when someone wins 11 points with margin. Win four out of 7 sets (similar: tennis, badminton, volleyball).
- **Boxing:** stop when essentially obvious evidence appears.
- **Quidditch:** game ends immediately when the Snitch is caught.

## Motivation: knock-out tournament design

In these games sometimes we think later stage games are more important:

- **Table tennis/badminton.** Early rounds are often best of 3 sets, while semifinals/finals are 5 or 7 sets.
- **Boxing.** Title fights are scheduled for 12 rounds, while non titled 4-10 rounds

**Does this make intuitive sense:?** The final determines the champion, while earlier rounds are numerous and each is less “informative” for the final winner.

**Statistical viewpoint:** longer matches reduce randomness and make the stronger player more likely to advance.

In this talk we study how to *optimally allocate match lengths* in a  $2^n$ -player knock-out tournament. We fix a target tournament accuracy

$$\mathbb{P}(\text{best player wins the tournament}) \geq 1 - \eta.$$

**Two design questions:**

- ① **Sequential vs. fixed length:** How much can we reduce the *average time per game* by using sequential stopping rules instead of fixed-length matches?
- ② **Round allocation:** Should later stage games be longer? If yes, what is the *optimal round-dependent schedule*?

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To answer the first question, we focus on the classical testing problem for a Brownian motion with unknown drift.

For each match we compare two testing methods:

- **Sequential testing:** where the match ends when evidence is strong enough.
- **Fixed-length testing:** a pre-determined sample size (fixed match length), chosen to achieve the same error level.

- The behavior of the players are often associated with some uncertainty.
- When a match is played, we can observe its “behavior” as a continuous process.
- We model the match behavior of Player 1 and 2 as a Brownian motion with drift:

$$X_t = \theta t + W_t,$$

where the unknown drift is unobservable:

$$\theta = \begin{cases} \frac{1}{2}, & \text{if Player 1 is better than Player 2,} \\ -\frac{1}{2}, & \text{if Player 2 is better than Player 1,} \end{cases}$$

- We thus test against the hypotheses  $H_+$  and  $H_-$ , where

$$H_+ : \theta = \frac{1}{2}, \quad H_- : \theta = -\frac{1}{2}.$$

- In a **fixed-size** sample test, one observes the process  $X$  over a time-interval  $[0, T]$ , where  $T$  is a constant chosen **before the sample is collected**.
- In a **sequential** test, choose a stopping time  $\tau$  once sufficient evidence is obtained.
- Naturally, the average sample size in a sequential test ( $\mathbb{E}[\tau]$ ) is smaller than the fixed sample size  $T$ .
- We study the average relative reduction  $\frac{T - \mathbb{E}[\tau]}{T}$  as a function of the power of the test.
- Denote  $1 - \eta \in (\frac{1}{2}, 1)$  the power of a test, so that  $\eta \in (0, \frac{1}{2})$  is the maximal probability with which the wrong hypothesis is accepted.
- We require that

$$\mathbb{P}_1(d = 1) \geq 1 - \eta \quad \& \quad \mathbb{P}_{-1}(d = -1) \geq 1 - \eta.$$

- By symmetry, an optimal decision at  $T$  is given by

$$d = \begin{cases} \frac{1}{2} & \text{if } X_T \geq 0 \\ -\frac{1}{2} & \text{if } X_T < 0. \end{cases}$$

- Therefore, we choose  $T \geq T_\eta$ , where  $T_\eta$  is defined by

$$\mathbb{P}_1(X_{T_\eta} \leq 0) = \mathbb{P}_{-1}(X_{T_\eta} \geq 0) = \Phi\left(-\sqrt{T_\eta}/2\right) = \eta,$$

where  $\Phi$  and  $\varphi$  are the CDF and pdf of  $N(0, 1)$ .

- Consequently,

$$T_\eta = 4 \left(\Phi^{-1}(\eta)\right)^2.$$

- In sequential experiments, the process is observed until an **information level** is achieved.
- Define the posterior probability process

$$\Pi_t := \mathbb{P}(\theta = \frac{1}{2} | \mathcal{F}_t)$$

- The process is observed until  $\tau := \inf\{t \geq 0 : \Pi_t \notin (b, 1-b)\}$  for some  $b \in (0, \frac{1}{2})$ .
- Declare  $\theta = \frac{1}{2}$  if  $\Pi_\tau = 1 - b$  and  $\theta = -\frac{1}{2}$  otherwise.
- We can show that

$$\mathbb{P}(\theta = d) = 1 - b.$$

- Therefore, to have the same power of the test, we let  $b = 1 - \eta$ .

## Sequential tests

The expected value of observation time  $\mathbb{E}[\tau]$  is characterized through finding the unique solution of the boundary value problem

$$\begin{cases} \frac{1}{2}\pi^2(1-\pi)^2u_{\pi\pi} + 1 = 0, & \pi \in (\eta, 1-\eta), \\ u(\eta) = u(1-\eta) = 0, \end{cases}$$

with  $\mathbb{E}[\tau] = u(\frac{1}{2}) = -2\psi(\eta)$ , where  $\psi(x) := (1-2x)\ln\frac{x}{1-x}$ .

### Remark 1

By Wald, there is one test that simultaneously minimizes  $\mathbb{E}_{-1}[\tau]$  and  $\mathbb{E}_1[\tau]$  over all tests with power  $1-\eta$ . The stopping time in that test is given by

$$\tau_\eta := \inf \left\{ t \geq 0 : X_t \notin \left( \log \frac{\eta}{1-\eta}, \log \frac{1-\eta}{\eta} \right) \right\}.$$

The average sample size then satisfies

$$\mathbb{E}_{-1}[\tau_\eta] = \mathbb{E}_1[\tau_\eta] = -2\psi(\eta).$$

- We introduce the function  $f : (0, \frac{1}{2}) \rightarrow [0, 1]$  given by

$$f(\eta) := \frac{\mathbb{E}_1[\tau_\eta]}{T_\eta} = \frac{-\psi(\eta)}{2(\Phi^{-1}(\eta))^2}. \quad (1)$$

- $f(\eta) < 1$ , and  $1 - f(\eta)$  represents the average sample size reduction by using sequential tests.
- We provide precise bounds on  $f(\eta)$ .

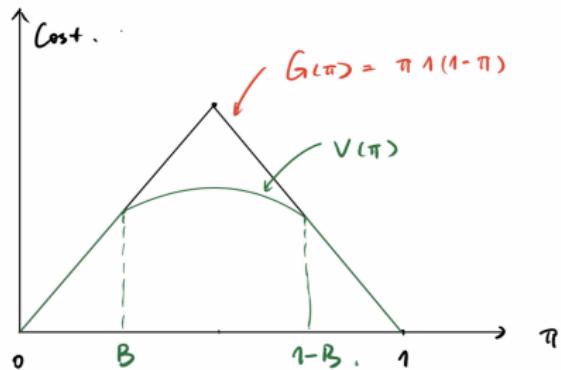
## Remark 1. Equivalency to Shiryaev's testing problem

Recall the Bayesian formulation with constant cost per unit time:

$$V = \inf_{\tau, d} \{ \mathbb{P}(d = 0, \Theta = 1) + \mathbb{P}(d = 1, \Theta = 0) + c\mathbb{E}[\tau] \}. \quad (2)$$

Problem (2) can be written as

$$V(\pi) = \inf_{\tau} \mathbb{E}_{\pi} [c\tau + \Pi_{\tau} \wedge (1 - \Pi_{\tau})]$$



- The formulation with  $\eta$  is equivalent to the classical testing problem with  $c$  via a Lagrange formulation.
- In other words, the optimized Lagrange multiplier corresponds to  $\frac{1}{c}$ .

## Remark 2. dependence on signal-to-noise ratio

- If the observation process is given by  $X_t = \theta\mu t + W_t$ , then the signal-to-noise ratio is  $\mu$ .
- By Brownian scaling, the process

$$\tilde{X}_t := \mu X_{t/\mu^2} = \mu \left( \theta \frac{t}{\mu^2} + W_{t/\mu^2} \right) = \theta t + \tilde{W}_t,$$

where  $\tilde{W}_t := \mu W_{t/\mu^2}$  is a standard Brownian motion.

- Then  $T_{\mu,\eta}$  needed to achieve a certain precision  $1 - \eta$  when the signal-to-noise ratio is  $\mu$  satisfies

$$T_{\mu,\eta} = \frac{1}{\mu^2} T_{1,\eta} = \frac{1}{\mu^2} T_\eta.$$

- Similarly,  $\mathbb{E}[\tau_{\mu,\eta}] = \frac{1}{\mu^2} E[\tau_{1,\eta}] = \frac{1}{\mu^2} E[\tau_\eta]$ .
- Both scale inversely with  $\mu^2$ , but the average reduction ratio  $1 - f(\eta)$  is independent of  $\mu$ . It suffices to consider  $\mu = \frac{1}{2}$ .

The average reduction ratio is increasing in the error  $\eta$ .

## Theorem 2

The function  $f$  is increasing on  $(0, \frac{1}{2})$ , with  $f(0+) = \frac{1}{4}$  and  $f(\frac{1}{2}-) = \frac{2}{\pi}$ .

**Remark.** The minimal reduction is 36% and the maximal reduction is 75%.  $f(0+) = \frac{1}{4}$  is well known [cf. Aivazjan('59), Shiyaev('78)], and numerically [cf. Eisenberg('91)] that convergence to 75% is rather slow, and that typical reductions range between 50% and 60%.

The monotonicity is perhaps obvious but new.

Table: Relative efficiency of SPRT.

| Power $1 - \eta$ | $f(\eta)$ | Average reduction |
|------------------|-----------|-------------------|
| 0.80             | 0.5871    | 41.29%            |
| 0.90             | 0.5351    | 46.49%            |
| 0.95             | 0.4897    | 51.03%            |
| 0.99             | 0.4160    | 58.40%            |
| 0.999            | 0.3609    | 63.91%            |

The average reduction ratio is increasing in the error  $\eta$ .

### Proposition 3

We have  $f(\eta) = \frac{1}{4} - \frac{\ln(-\ln \eta)}{8 \ln \eta} + o\left(\frac{\ln(-\ln \eta)}{\ln \eta}\right)$  as  $\eta \rightarrow 0$ .

**Remark.**  $f$  approaches its limit at  $\eta = \frac{1}{2}$  quadratically.

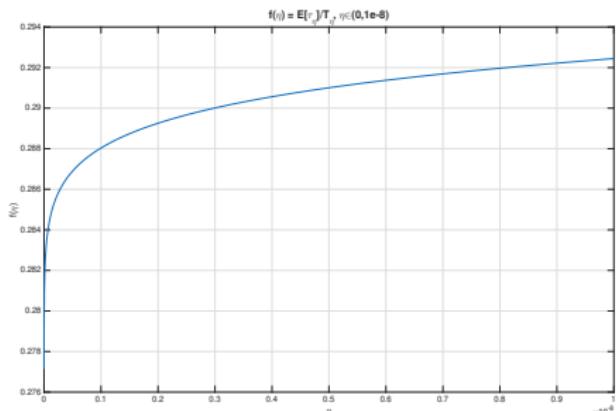
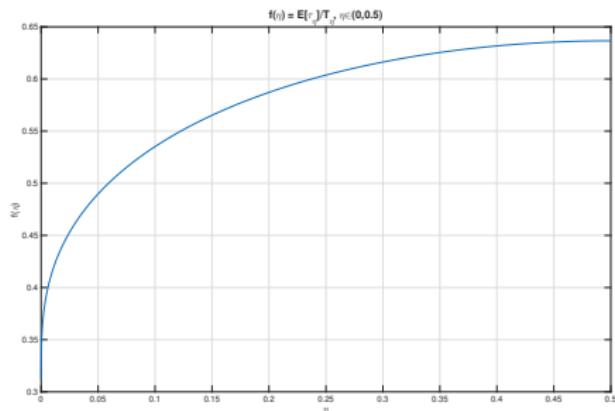


Figure: The function  $f(\eta)$  for  $\eta \in (0, \frac{1}{2})$  (left), and  $f(\eta)$  for  $\eta \in (0, 10^{-8})$  (right).

To reach a 70%, 72%, 74% efficiency reduction, we need error probabilities  $\eta$  to be  $2 \times 10^{-7}, 2 \times 10^{-12}, 2 \times 10^{-40}$ , respectively.

# Outline

1 Motivation

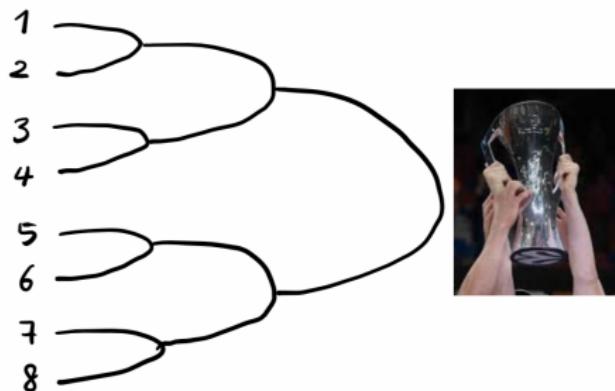
2 Single match setting

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# Problem formulation

There are many different forms of tournaments. Of the knockout types:

- Football, table tennis,...



The champion!

- Children game, sorting:



The "king of the hill"

## Problem formulation

Consider a game of  $2^n$  players (of football type). Want to minimize the observation cost (time length), subject to some probability error.

### Assumptions:

- There is a distinct rank.
- Each match is a Brownian motion

$$X_t^{ij} = \Theta^{ij} t + W_t^{ij},$$

where

$$\Theta^{ij} = \begin{cases} \frac{1}{2}, & \text{if } i \text{ is better than } j, \\ -\frac{1}{2}, & \text{if } j \text{ is better than } i, \end{cases}$$

- Initially, we have a uniform prior distribution on  $n!$  configurations. e.g., with  $n = 2$ :

$$\left\{ \begin{pmatrix} 1 \\ 2 \\ 3 \\ 4 \end{pmatrix}, \begin{pmatrix} 1 \\ 2 \\ 4 \\ 3 \end{pmatrix}, \dots, \begin{pmatrix} 4 \\ 3 \\ 2 \\ 1 \end{pmatrix} \right\}, \quad \text{(24 configurations)}$$

## Matches of mixed lengths

- Each match in stage  $k$  is played until a pre-determined time  $T_k$  is reached.
- The winner is Player  $i$  if  $X_{T_k}^{ij} \geq 0$ , and Player  $j$  otherwise.
- We assume every game is **equally costly**.
- Fix  $\eta \in (0, 2^{-n})$ , we consider

$$\inf_{\{T_k\}_{k=1}^n} \sum_{k=1}^n 2^{n-k} T_k$$

subject to the constraint

$$\mathbb{P}(\text{the best player wins the tournament}) \geq 1 - \eta.$$

- Since a better player wins a match of length  $T$  with probability  $\Phi(\frac{\sqrt{T}}{2})$ , this is equivalent to

$$\inf_{\{T_k\}_{k=1}^n} \sum_{k=1}^n 2^{n-k} T_k \quad \text{subject to} \quad \prod_{k=1}^n \Phi\left(\frac{\sqrt{T_k}}{2}\right) \geq 1 - \eta. \quad (3)$$

## Matches of fixed lengths

- Define the Lagrangian

$$\mathcal{L}(\{T_k\}_{k=1}^n, \lambda) := \sum_{k=1}^n 2^{n-k} T_k - \lambda \left( \sum_{k=1}^n a(T_k) - \log(1 - \eta) \right)$$

with  $a(T) := \log(\Phi(\frac{\sqrt{T}}{2}))$ .

- $a$  is increasing and strictly concave.
- For the  $k$ th round, take the first order condition:  $\frac{\partial \mathcal{L}}{\partial T_k} = 2^{n-k} - a'(T_k) = 0$ .

### Theorem 4

Let  $\eta \in (0, 2^{-n})$  be given. Then there is a unique solution

$(T_1, \dots, T_n) = (T_1^*, \dots, T_n^*)$  with  $T_k^* > 0$  for  $k = 1, \dots, n$  to the system

$$\begin{cases} \frac{a'(T_1)}{2^{n-1}} = \frac{a'(T_2)}{2^{n-2}} = \dots = \frac{a'(T_n)}{2} \\ \sum_{k=1}^n a(T_k^*) = \log(1 - \eta), \end{cases}$$

and  $(T_1^*, \dots, T_n^*)$  solves the optimization problem (3). Moreover,  
 $T_1^* < T_2^* < \dots < T_n^*$ .

## Sequential matches

- Each match in stage  $k$  is played until an **information level** is achieved.
- Define the posterior probability process

$$\Pi_t^{ij} := \mathbb{P}(\theta_{ij} = \frac{1}{2} | \mathcal{F}_t^{ij})$$

- Important assumption: **symmetry**.
- The games in the same round uses the same strategy. i.e., the match is thus played until  $\tau_k^{ij} := \tau^{ij}(b_k) := \inf\{t \geq 0 : \Pi_t^{ij} \notin (b_k, 1 - b_k)\}$ ,
- Declare the winner to be Player  $i$  if  $\Pi_{\tau_k}^{ij} = 1 - b_k$  and Player  $j$  if  $\Pi_{\tau_k}^{ij} = b_k$ .
- Denote the expected value of observation time  $\tau_k$

$$\mathcal{T}(b_k) := \mathbb{E}[\tau_k] = -2\psi(b_k).$$

- Consequently, the constrained sequential tournament design problem becomes

$$\inf_{\mathbf{b} \in (0, \frac{1}{2})^n} \sum_{k=1}^n 2^{n-k} \mathcal{T}(b_k) \quad \text{subject to} \quad \prod_{k=1}^n (1 - b_k) \geq 1 - \eta. \quad (4)$$

## Sequential matches

- Similar to the fixed-time case, define

$$\mathcal{L}(\mathbf{b}, \lambda) := \sum_{k=1}^n 2^{n-k} \mathcal{T}(b_k) - \lambda \left( \sum_{k=1}^n \log(1 - b_k) - \log(1 - \eta) \right).$$

- The first order condition gives

$$\frac{\partial \mathcal{L}}{\partial b_k} = 2^{n-k} \mathcal{T}'(b_k) + \lambda \frac{1}{1 - b_k} = 0.$$

- $(1 - b) \mathcal{T}'(b)$  is strictly negative and increasing.

### Theorem 5

Let  $\eta \in (0, 2^{-n})$  be given. Then there is a unique solution  $(b_1, \dots, b_n) = (b_1^*, \dots, b_n^*)$  with  $b_k^* \in (0, 1/2)$  for  $k = 1, \dots, n$  to the system

$$\begin{cases} 2^{n-k}(1 - b_k) \mathcal{T}'(b_k) = 2^{n-l}(1 - b_l) \mathcal{T}'(b_l), \\ \sum_{k=1}^n \log(1 - b_k) = \log(1 - \eta). \end{cases}$$

Moreover,  $(b_1^*, \dots, b_n^*)$  solves the optimization problem (4), and  $b_1^* \geq b_2^* \geq \dots \geq b_n^*$ .

- ① **Should always play sequential.**
- ② **If we have to play fixed-length, let's observe more at later rounds.**

$$T_1^* \leq T_2^* \leq \cdots \leq T_n^*.$$

- ③ **When we play sequential, should also observe more at later rounds.**

$$b_1^* \geq b_2^* \geq \cdots \geq b_n^*, \quad \mathbb{E}[\tau_1^*] \leq \mathbb{E}[\tau_2^*] \leq \cdots \leq \mathbb{E}[\tau_n^*].$$

**Remark.** Cost of each game affects the monotonicity.

# Thank you!